

## Addendum to: Entropic Value-at-Risk: A New Coherent Risk Measure

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**Abstract** This short addendum consists of two sections. The first provides proofs that were omitted in Ahmadi-Javid (J. Optim. Theory Appl., 2012) for the sake of brevity, and also demonstrates that the dual representation of the entropic value-at-risk, which is given in Ahmadi-Javid (J. Optim. Theory Appl., 2012) for the case of bounded random variables, holds for all random variables whose moment-generating functions exist everywhere. The second section provides a few corrections.

**Keywords** Chernoff inequality · Coherent risk measure · Conditional value-at-risk (CVaR) · Convex optimization · Cumulant-generating function · Duality · Entropic value-at-risk (EVaR) · g-entropic risk measure · Moment-generating function · Relative entropy · Stochastic optimization · Stochastic programming · Value-at-risk (VaR)

### 1 Supplementary Proofs

In this section, we begin by providing detailed proofs for some of the statements made in [1], which may be helpful to readers who are less familiar with convex optimization. Then we discuss the dual representation of the entropic value-at-risk (EVaR) for any random variable whose moment-generating function exists everywhere.

The following lemma proves the convexity of cumulant-generating functions, which is used in the proof of Lemma 3.1 of [1].

**Lemma 1.1** *The cumulant-generating function  $\ln M_X(t)$  is convex in  $X$ .*

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*Proof* Without any loss of generality, we can set  $t = 1$ . It now suffices to show that, for any two random variables  $X$  and  $Y$  with finite  $M_X(1)$ ,  $M_Y(1)$  and any  $\lambda \in [0, 1]$ ,

$$E(e^{\lambda X + (1-\lambda)Y}) \leq E(e^X)^\lambda E(e^Y)^{1-\lambda}.$$

Defining  $W := e^X/E(e^X)$  and  $V := e^Y/E(e^Y)$ , this inequality is equivalent to  $E(W^\lambda V^{1-\lambda}) \leq 1$  that immediately follows from the inequality of weighted arithmetic and geometric means for the non-negative random variables  $W$  and  $V$ , i.e.,  $W^\lambda V^{1-\lambda} \leq \lambda W + (1 - \lambda)V$ . □

The following lemma proves a statement used in the proof of Theorem 3.1 of [1].

**Lemma 1.2** *The function  $\inf_{t>0}\{\kappa_\alpha(X, t)\}$  is convex in  $X$  for all  $\alpha \in ]0, 1]$ .*

*Proof* For all  $\varepsilon > 0$ , and  $X, Y \in \mathbf{L}_{M^+}$ , the continuity of  $\kappa_\alpha(X, t)$  in  $t > 0$  implies the existence of  $t_1, t_2 > 0$  such that

$$\kappa_\alpha(X, t_1) \leq \inf_{t>0}\{\kappa_\alpha(X, t)\} + \varepsilon, \quad \kappa_\alpha(Y, t_2) \leq \inf_{t>0}\{\kappa_\alpha(Y, t)\} + \varepsilon.$$

Since  $\kappa_\alpha(X, t)$  is convex in  $(X, t)$  from Lemma 3.1 of [1], we further find that, for all  $\lambda \in [0, 1]$ ,

$$\begin{aligned} \inf_{t>0}\{\kappa_\alpha(\lambda X + (1 - \lambda)Y, t)\} &\leq \kappa_\alpha(\lambda X + (1 - \lambda)Y, \lambda t_1 + (1 - \lambda)t_2) \\ &\leq \lambda \kappa_\alpha(X, t_1) + (1 - \lambda)\kappa_\alpha(Y, t_2) \leq \lambda \inf_{t>0}\{\kappa_\alpha(X, t)\} + (1 - \lambda) \inf_{t>0}\{\kappa_\alpha(Y, t)\} + \varepsilon. \end{aligned}$$

As this holds for all  $\varepsilon > 0$ , the proof is done by taking the limit  $\varepsilon \downarrow 0$ . □

The following lemma proves the validity of the last identities used in the proofs of Theorems 3.3 and 5.1 of [1].

**Lemma 1.3** *Let  $g : \mathbb{R} \rightarrow [0, \infty]$  be a non-negative closed convex function with  $g(1) = 0$  and  $\text{dom}g = [\gamma_1, \gamma_2]$  where  $\gamma_1 < 1 < \gamma_2$  and  $\gamma_1, \gamma_2 \in [-\infty, \infty]$ . Then the following identity holds for any  $\beta > 0$*

$$\inf_{t>0}\left\{\sup_{Q \ll P}\{E_Q(X) + t(\beta - H_g(P, Q))\}\right\} = \sup_{Q \ll P, H_g(P, Q) \leq \beta}\{E_Q(X)\}.$$

*Proof* By denoting  $Y = \frac{dQ}{dP}$ , which is a non-negative random variable with the mean equal to 1, the above identity can be rewritten as

$$\sup_{t \geq 0}\{L(t)\} = \inf_{Y \in S, E_P(g(Y)) \leq \beta}\{-E_P(XY)\},$$

where  $L(t) = \inf_{Y \in S}\{-E_Q(XY) + t(E_P(g(Y)) - \beta)\}$  is the Lagrangian associated with the optimization problem on the right-hand side, and  $S = \{Y \in \mathbf{L}_1 : E_P(Y) = 1, \max\{\gamma_1, 0\} \leq Y \leq \gamma_2 \text{ a.e.}\}$ . Hence, it suffices to show that the optimal duality gap for the right-hand side optimization problem is zero. This is possible by showing that the generalized Slater’s constraint qualification [2] holds for this problem, i.e., that there exists  $\hat{Y} \in \mathbf{L}_1$  satisfying  $E_P(\hat{Y}) = 1, \max\{\gamma_1, 0\} < \hat{Y} < \gamma_2 \text{ a.e.}, E_P(g(\hat{Y})) < \beta$ .

As we assumed  $\gamma_1 < 1 < \gamma_2$ , the solution  $\widehat{Y} = 1$  a.e. fulfills these conditions, and so the proof is complete.  $\square$

**Remark 1.1** We can also show the validity of the identity in Lemma 1.3 for the case where  $g$  is not always non-negative over its domain. If  $g$  is not non-negative, then, assuming that  $c(x - 1)$  is a supporting hyperplane to the epigraph of  $g$  at the point  $(1, 0)$ , it is sufficient to replace  $g$  with the non-negative function  $\tilde{g}(x) = g(x) - c(x - 1)$ , for which we have  $H_{\tilde{g}}(P, Q) = H_g(P, Q)$ . Moreover, if  $\gamma_1 = 1$  or  $\gamma_2 = 1$ , then the validity of the identity in Lemma 1.3 is clear, because the constraint  $E_P(\frac{dQ}{dP}) = 1$  implies  $\frac{dQ}{dP} = 1$  a.e. or, equivalently,  $Q = P$ . Finally, note that the proof for the case  $\beta = 0$  is also straightforward, given that  $\gamma_1 < 1 < \gamma_2$  and  $g$  is non-negative. Indeed, for this case, the constraint  $H_g(P, Q) \leq 0$  is feasible if and only if  $g$  is zero over the interval  $\text{dom } g = [\gamma_1, \gamma_2]$ , which implies that the constraint  $H_g(P, Q) \leq 0$  is redundant since it is equivalent to  $\gamma_1 \leq \frac{dQ}{dP} \leq \gamma_2$  a.e. which already exists in the set  $S$ .

Here we give the detailed proof of Proposition 4.2 of [1].

*Proof of Proposition 4.2 of [1]* To use Theorem 4.1 of [1], we first need to reformulate problem (8) of [1] by using the CVaR representation given in (2) of [1] and adding the additional constraint  $lC \leq -t \leq uC$ . The resulting problem can be rewritten in the form of problem (6) of [1] as follows:

$$\min_{w \in \mathbf{W}, lC \leq -t \leq uC} \left\{ t + \alpha^{-1} E \left[ - \sum_{i=1}^n w_i R_i - t \right]_+ \right\} = \min_{\mathbf{x} \in \mathbf{X}} E(F(\mathbf{x}, \xi)),$$

where  $\mathbf{x} = (w^T, t)^T$ ,  $\xi = \mathbf{R}$ ,  $\mathbf{X} = \mathbf{W} \times [-uC, -lC]$  and  $F(\mathbf{x}, \xi) = t + \alpha^{-1} [- \sum_{i=1}^n w_i R_i - t ]_+$ . Then we need to find  $D$  and  $L$ . In this case,  $D = \sup_{\mathbf{x}, \mathbf{x}' \in \mathbf{X}} \|\mathbf{x} - \mathbf{x}'\| = C\sqrt{2 + B^2}$ , where the maximum value is attained for  $\mathbf{x} = (C, 0, \dots, 0, -uC)^T$  and  $\mathbf{x}' = (0, C, \dots, 0, -lC)^T$ , or other similar pairs of points. To determine the Lipschitz constant  $L$ , we have

$$\begin{aligned} |F(\mathbf{x}, \mathbf{z}) - F(\mathbf{x}', \mathbf{z})| &\leq |t - t'| + \alpha^{-1} |[-\mathbf{z}^T \mathbf{w} - t]_+ - [-\mathbf{z}^T \mathbf{w}' - t']_+| \\ &\leq |t - t'| + \alpha^{-1} |[-\mathbf{z}^T (\mathbf{w} - \mathbf{w}') - (t - t')]_+| \\ &\leq |t - t'| + \alpha^{-1} |\mathbf{z}^T (\mathbf{w} - \mathbf{w}') + (t - t')| \\ &= |t - t'| + \alpha^{-1} |(\mathbf{z}^T, 1)(\mathbf{x} - \mathbf{x}')| \leq (1 + \alpha^{-1} \|(\mathbf{z}^T, 1)^T\|) \|\mathbf{x} - \mathbf{x}'\| \\ &\leq (1 + \alpha^{-1} \sqrt{n \max\{u^2, l^2\} + 1}) \|\mathbf{x} - \mathbf{x}'\|, \end{aligned}$$

where, in the last two inequalities, we used the Cauchy–Schwarz inequality and the fact that  $\mathbf{z} \in S_\xi = S_R = [l, u]^n$ . This shows that  $L = 1 + \alpha^{-1} \sqrt{n \max\{u^2, l^2\} + 1}$ .  $\square$

The next theorem proves that the dual representation in Theorem 3.3 of [1], which is given for bounded random variables, also holds for  $X \in \mathbf{L}_M$ .

**Theorem 1.1** (Dual representation of the EVaR) *For  $X \in \mathbf{L}_M$  and any  $\alpha \in ]0, 1]$*

$$\text{EVaR}_{1-\alpha}(X) = \sup_{Q \in \mathfrak{S}} E_Q(X),$$

where  $\mathfrak{S} = \{Q \ll P : D_{KL}(Q \parallel P) \leq -\ln \alpha\}$ .

*Proof* By virtue of the result given in Sect. 5.4 of [3], for any  $X \in \mathbf{L}_M$ , we have

$$\ln E_P(e^X) = \sup_{Q \in \mathfrak{S}'} \{E_Q(X) - D_{KL}(Q \parallel P)\},$$

where

$$\mathfrak{S}' = \left\{ Q \ll P : \exists c > 0 : E_P \left( h \left( c \frac{dQ}{dP} \right) \right) < \infty \right\}$$

$$\text{with } h(x) = \begin{cases} 0, & 0 \leq x < 1, \\ x \ln x - x + 1, & 1 \leq x. \end{cases}$$

Hence, similarly as in the proofs of Lemma 1.3 and Theorem 3.3 of [1], and together with Remark 1.1, we can show

$$\text{EVaR}_{1-\alpha}(X) = \sup_{Q \in \mathfrak{S}''} E_Q(X),$$

where  $\mathfrak{S}'' = \{Q \ll P : D_{KL}(Q \parallel P) \leq -\ln \alpha, \exists c > 0 : E_P(h(c \frac{dQ}{dP})) < \infty\}$ . However, one can see that the constraint  $D_{KL}(Q \parallel P) \leq -\ln \alpha$  implies  $E_P(h(c \frac{dQ}{dP})) < \infty$  with  $c = 1$ , because the relative entropy can be rewritten as  $D_{KL}(Q \parallel P) = E_P(e(\frac{dQ}{dP}))$  with  $e(x) = x \ln x - x + 1, x > 0$ . Hence, the two sets  $\mathfrak{S}$  and  $\mathfrak{S}''$  are actually identical. This completes the proof. □

## 2 Corrections

This section corrects a few errors found in [1]. Since the numbering system of [1] was changed for publication, there are now two lemmas numbered 3.1: one precedes Theorem 3.1 of [1] and the other precedes Theorem 3.3 of [1]. The former of these lemmas was only used in the proof of Theorem 3.1 of [1], but the latter was used both in the proof of Theorem 3.3 of [1], and in the statement of Lemma 5.1 of [1]. Furthermore, at the end of the proof of the former, there are a few unnecessary sections that should be removed. Here is the corrected proof.

*Proof of Lemma 3.1 of [1]* We must show that, for all  $\lambda \in [0, 1], X, Y \in \mathbf{L}_M^+$  and  $t_1, t_2 > 0$ ,

$$\lambda \kappa_\alpha(X, t_1) + (1 - \lambda) \kappa_\alpha(Y, t_2) \geq \kappa_\alpha(\lambda X + (1 - \lambda)Y, \lambda t_1 + (1 - \lambda)t_2),$$

which is equivalent to

$$\lambda t_1 \ln M_X(t_1^{-1}) + (1 - \lambda)t_2 \ln M_Y(t_2^{-1})$$

$$\geq (\lambda t_1 + (1 - \lambda)t_2) \ln M_{\lambda X + (1 - \lambda)Y}((\lambda t_1 + (1 - \lambda)t_2)^{-1}).$$

Denoting  $t = \lambda t_1 + (1 - \lambda)t_2$  and  $w = \lambda t_1/t$ , the left-hand side of the above inequality can be expressed as

$$t(w \ln M_X(t_1^{-1}) + (1 - w) \ln M_Y(t_2^{-1})).$$

Then by using the known fact that the cumulant-generating function is convex, it yields

$$\begin{aligned} t(w \ln M_X(t_1^{-1}) + (1 - w) \ln M_Y(t_2^{-1})) &\geq t \ln E(e^{wXt_1^{-1} + (1-w)Yt_2^{-1}}) \\ &= t \ln E(e^{\lambda X t_1^{-1} + (1-\lambda)Y t_2^{-1}}) \\ &= (\lambda t_1 + (1 - \lambda)t_2) \ln M_{\lambda X + (1-\lambda)Y}((\lambda t_1 + (1 - \lambda)t_2)^{-1}). \end{aligned}$$

This completes the proof.  $\square$

In the first sentence, coming after the proof of Proposition 4.2 of [1], the term “*the feasible set of problem (8)*” must be replaced by “*the feasible set of the problem obtained from problem (8) by using equation (2)*.” Note that the problem obtained from problem (8) of [1] by using equation (2) of [1] is as follows:

$$\min_{w \in \mathbf{W}, t \in \mathbb{R}} \left\{ t + \alpha^{-1} \mathbb{E} \left[ - \sum_{i=1}^n w_i R_i - t \right]_+ \right\}.$$

## References

1. Ahmadi-Javid, A.: Entropic value-at-risk: A new coherent risk measure. J. Optim. Theory Appl. (2012), this issue
2. Jeyakumar, V., Wolkowicz, H.: Generalizations of Slater’s constraint qualification for infinite convex programs. Math. Program., Ser. B **57**, 85–101 (1992)
3. Cheridito, P., Li, T.: Risk measures on Orlicz hearts. Math. Finance **19**, 189–214 (2009)